

ANALYSIS OF NETWORKS WITH TIME-DEPENDENT TRANSITION PROBABILITIES AND MESSAGES BYPASS BETWEEN THE QUEUING SYSTEMS

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Abstract. The object of investigation is an open exponential network with a messages bypass of systems in transient behavior. The purpose of the research is to find stationary probabilities of states and the average characteristics of the network when the transition probabilities between the messages and bypass systems of the network, parameters of the incoming flow of messages and services are time-dependent. To find the state probabilities and the characteristics of a network is used the apparatus for the multivariate generating functions. The examples are calculated on a computer.

1. General information

The results of research of the above networks in stationary behavior are given in [1-3]. In [4] investigated the network in a transition behavior, but when the probability messages bypass between systems, networks do not depend on network status and time. To find the state probabilities of the network in a transition mode, we used the method of multidimensional generating functions, which was previously used to find the state probabilities of the other networks [5, 6]. In this paper we consider the case when the transition probabilities and messages bypass between the network systems depend on time.

2. Formulation of the problem

Consider an open exponential QN of arbitrary structure consisting of n QS, enumerated by numbers from 1 to n . Messages have a chance to join the queue, and with an additional probability to move immediately in accordance with the transition probability matrix of the other QS, or leave the network. The probability of joining the QS depends on the state of the QS and the number of QS with which the messages are sent to this QS. It is assumed that the incoming flow of applications to the network is simple. The results of studies of such networks in the steady state are given in [1-4]. This paper describes a method of finding the time-dependent state probabilities of the network of such a network in the transient state.

Let m_i - number of identical service lines in the QS S_i , I_i - a vector of dimension n , consisting of zeros except the i -th component, which is equal to 1, $i = \overline{1, n}$; p_{ij} - the transition probability of the message after service in the system S_i into the system S_j , $i, j = \overline{0, n}$, we assume the system S_0 is the external environment. Let us consider the case when the parameters of the incoming flow of messages and services depend on time, i.e. the time interval $[t, t + \Delta t)$ in the network receives an message with a probability $\lambda(t)\Delta t + o(\Delta t)$, and if at the time t of service on the line i -th QS is located an message, at the range $[t, t + \Delta t)$ of its services will end with a probability $\mu_i(t)\Delta t + o(\Delta t)$, $i = \overline{1, n}$. The message is sent to the i -th QS with probability p_{0i} , $\sum_{i=1}^n p_{0i} = 1$. The message sent to this QS from the external environment at moment time t , with a probability $f^{(i)}(k, t)$, when the network is in a state (k, t) , joins the queue, and the probability $1 - f^{(i)}(k, t)$ is not attached to the queue, regardless of handled (i.e., it's time of service with a probability of 1 is equal to zero). If the message has been served in the i -th QS, it is likely to be sent immediately to the j -th QS with probability p_{ij} , and leaves the QN with the probability p_{in} , $\sum_{j=0}^n p_{ij} = 1$, $i = \overline{1, n}$.

Let $k(t) = (k, t) = (k_1, k_2, \dots, k_n, t)$ - the state vector of the network, where k_i - the number of messages at the moment t in the system S_i , $i = \overline{1, n}$; $\varphi_i(k, t)$ - the conditional probability that the message is delivered to the i -th QS at time t , when the network is in a state (k, t) , will not be serviced by any of the QS; $\psi_{ij}(k, t)$ - the conditional probability that the message is delivered to the i -th QS outside at time t , when the network is in state (k, t) , the first time, a service in j -th QS; $\alpha_i(k, t)$ - the conditional probability that the message, served in the i -th queuing system at time t , when the network is in a state (k, t) , will no longer be served in any of QS; $\beta_{ij}(k, t)$ - the conditional probability that the message, served in the i -th queuing system at time t , when the network is in state (k, t) for the first time then receive services in the j -th QS, $i, j = \overline{1, n}$.

According to the formula of total probability, we obtain:

$$\begin{aligned} \varphi_i(k, t) &= (1 - f^{(i)}(k, t)) \left(p_{in} + \sum_{j=1}^n p_{ij} \varphi_j(k, t) \right), \quad i = \overline{1, n}, \\ \psi_{ij}(k, t) &= f^{(i)}(k, t) \delta_{ij} + (1 - f^{(i)}(k, t)) \sum_{d=1}^n p_{id} \psi_{dj}(k, t), \quad i, j = \overline{1, n}, \end{aligned} \quad (1)$$

$$\alpha_i(k, t) = p_{i0} + \sum_{j=1}^n p_{ij} \varphi_j(k - I_i, t), \quad i = \overline{1, n}, \quad (2)$$

$$\beta_{ij}(k, t) = \sum_{l=1}^n p_{il} \psi_{lj}(k - I_i, t), \quad i, j = \overline{1, n}, \quad (3)$$

where δ_{ij} - the Kronecker delta. We have the equalities

$$\alpha_i(k, t) = 1 - \sum_{j=1}^n \beta_{ij}(k, t), \quad \varphi_i(k, t) + \sum_{j=1}^n \psi_{ij}(k, t) = 1, \quad i, j = \overline{1, n}.$$

From (1) and (3) we find

$$\psi_{ij}(k, t) = f^{(i)}(k, t) \delta_{ij} + (1 - f^{(i)}(k, t)) \beta_{ij}(k - I_i, t), \quad i, j = \overline{1, n}. \quad (4)$$

The probabilities of states of the network under consideration satisfy the difference-differential equations (DDE) is proved in [4]:

$$\begin{aligned} \frac{dP(k, t)}{dt} = & - \sum_{i=1}^n [\lambda(t) p_{0i} (1 - \varphi_i(k, t)) + \mu_i(t) (1 - \beta_{ii}(k, t)) \min(m_i, k)] P(k, t) + \\ & + \lambda(t) \sum_{i=1}^n \sum_{j=1}^n p_{0i} \psi_{ij}(k - I_j, t) u(k_j) P(k - I_j, t) + \\ & + \sum_{i=1}^n \mu_i(t) \min(m_i, k_i + 1) \alpha_i(k + I_i, t) P(k + I_i, t) + \\ & + \sum_{\substack{i, j=1 \\ i \neq j}}^n \mu_i(t) \min(m_i, k_i + 1) \beta_{ij}(k + I_i - I_j, t) u(k_j) P(k + I_i - I_j, t), \end{aligned} \quad (5)$$

where $u(x) = \begin{cases} 1, & x > 0 \\ 0, & x \leq 0 \end{cases}$ - is the Heaviside function.

3. Finding the state probabilities

Let $m_i = 1, \quad i = \overline{1, n}$, and suppose that all network system operating in high load mode, i.e. $k_i(t) > 0 \quad \forall t > 0, \quad i = \overline{1, n}$, then the system (5) takes the form

$$\begin{aligned} \frac{dP(k, t)}{dt} = & - \sum_{i=1}^n [\lambda(t) p_{0i} (1 - \varphi_i(k, t)) + \mu_i(t) (1 - \beta_{ii}(k, t))] P(k, t) + \\ & + \lambda(t) \sum_{i=1}^n \sum_{j=1}^n p_{0i} \psi_{ij}(k - I_j, t) P(k - I_j, t) + \sum_{i=1}^n \mu_i(t) \alpha_i(k + I_i, t) P(k + I_i, t) + \\ & + \sum_{\substack{i, j=1 \\ i \neq j}}^n \mu_i(t) \beta_{ij}(k + I_i - I_j, t) P(k + I_i - I_j, t). \end{aligned} \quad (6)$$

Note, that the number of equations in (6) is countable, when the network is open, and of course, when it is closed.

We denote by $\Psi_n(z, t)$, where $z = (z_1, z_2, \dots, z_n)$, n -dimensional generating function:

$$\Psi_n(z, t) = \sum_{k_1=0}^{\infty} \sum_{k_2=0}^{\infty} \dots \sum_{k_n=0}^{\infty} P(k_1, k_2, \dots, k_n, t) z_1^{k_1} z_2^{k_2} \dots z_n^{k_n} = \sum_{k_1=1}^{\infty} \sum_{k_2=1}^{\infty} \dots \sum_{k_n=1}^{\infty} P(k, t) \prod_{i=1}^n z_i^{k_i}, \quad (7)$$

the summation is taken over each k_i , from 1 to ∞ , $i = \overline{1, n}$, because the network is operating in a high load mode.

Consider the case where the conditional probabilities $\varphi_i(k, t)$, $\psi_{ij}(k, t)$, $\alpha_i(k, t)$, $\beta_{ij}(k, t)$ do not depend on the state of the network. The expression for the generating function (7) can be rewritten as

$$\begin{aligned} \Psi_n(z, t) = C_n(z) \exp \left\{ - \int_{t_0}^t \left[\sum_{i=1}^n \left[\lambda(t) p_{wi} (1 - \varphi_i(t)) + \mu_i(t) (1 - \beta_{ii}(t)) + \right. \right. \right. \\ \left. \left. \left. + \lambda(t) p_{oi} z_i \sum_{j=1}^n \psi_{ij}(t) + \mu_i(t) \frac{\alpha_i(t)}{z_i} - \mu_i(t) \sum_{j=1}^n \beta_{ij}(t) \frac{z_j}{z_i} \right] dt \right\}, \quad (8) \end{aligned}$$

where the function $C_n(z)$ was defined in the proof of Lemma 2 in paper [4], from the conditions, that at the initial time the network is able to $(x_1, x_2, \dots, x_n, 0)$, $x_i > 0$, $i = \overline{1, n}$, $P(x_1, x_2, \dots, x_n, 0) = 1$, $P(k_1, k_2, \dots, k_n, 0) = 0$, $\forall x_i \neq k_i$, $i = \overline{1, n}$:

$$\begin{aligned} C_n(z) = \exp \left\{ \Lambda(0) \sum_{i=1}^n p_{wi} - \sum_{i=1}^n p_{wi} \Phi_i(0) + \sum_{i=1}^n M_i(0) - \sum_{i=1}^n B_{ii}(0) + \right. \\ \left. + \sum_{i=1}^n p_{oi} z_i \sum_{j=1}^n Y_{ij}(0) + \sum_{i=1}^n \frac{1}{z_i} A_i(0) + \sum_{i=1}^n \sum_{j=1}^n \frac{z_j}{z_i} B_{ij}(0) \right\} \prod_{i=1}^n z_i^{x_i}. \quad (9) \end{aligned}$$

Let us introduce the following notations:

$$\begin{aligned} \Lambda(t) = \int \lambda(t) dt, \quad M_i(t) = \int \mu_i(t) dt, \quad \Phi_i(t) = \int \lambda(t) \varphi_i(t) dt, \quad Y_{ij}(t) = \int \lambda(t) \psi_{ij}(t) dt, \\ A_i(t) = \int \mu_i(t) \alpha_i(t) dt, \quad B_{ij}(t) = \int \mu_i(t) \beta_{ij}(t) dt. \quad (10) \end{aligned}$$

Using the notation (10), expression (8) can be rewritten as

$$\begin{aligned} \Psi_n(z, t) = C_n(z) \exp \left\{ - \left(\Lambda(t) \sum_{i=1}^n p_{wi} - \sum_{i=1}^n p_{wi} \Phi_i(t) + \sum_{i=1}^n M_i(t) - \sum_{i=1}^n B_{ii}(t) + \right. \right. \\ \left. \left. + \sum_{i=1}^n p_{oi} z_i \sum_{j=1}^n Y_{ij}(t) + \sum_{i=1}^n \frac{1}{z_i} A_i(t) + \sum_{i=1}^n \sum_{j=1}^n \frac{z_j}{z_i} B_{ij}(t) \right) \right\}. \quad (11) \end{aligned}$$

Then from (9) and (11) that

$$\begin{aligned} \Psi_n(z, t) = \exp \left\{ - \left((\Lambda(t) - \Lambda(0)) \sum_{i=1}^n p_{vi} - \sum_{i=1}^n p_{vi} (\Phi_i(t) - \Phi_i(0)) + \sum_{i=1}^n (M_i(t) - M_i(0)) - \right. \right. \\ \left. - \sum_{j=1}^n (B_{nj}(t) - B_{nj}(0)) + \sum_{i=1}^n p_{vi} z_i \sum_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) + \sum_{i=1}^n \frac{1}{z_i} (A_i(t) - A_i(0)) + \right. \\ \left. + \sum_{i=1}^n \sum_{j=1}^n \frac{z_j}{z_i} (B_{ij}(t) - B_{ij}(0)) \right) \left. \right\} \prod_{i=1}^n z_i^{v_i}. \end{aligned}$$

This expression can be rewritten as

$$\begin{aligned} \Psi_n(z, t) = a_0(t) \exp \left\{ \sum_{i=1}^n p_{vi} z_i \sum_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right\} \times \\ \times \exp \left\{ \sum_{i=1}^n \frac{1}{z_i} (A_i(t) - A_i(0)) \right\} \exp \left\{ \sum_{i=1}^n \sum_{j=1}^n \frac{z_j}{z_i} (B_{ij}(t) - B_{ij}(0)) \right\} \prod_{i=1}^n z_i^{v_i} \\ a_0(t) = \exp \left\{ - \sum_{i=1}^n [(\Lambda(t) - \Lambda(0)) - (\Phi_i(t) - \Phi_i(0))] p_{vi} + \right. \\ \left. + (M_i(t) - M_i(0)) - (B_{ni}(t) - B_{ni}(0)) \right\}. \end{aligned} \quad (12)$$

Transform (12) to a form suitable for finding the state probabilities of the network, expanding its member exponential in a Maclaurin series. From (7) and (12) that

$$\Psi_n(z, t) = a_0(t) a_1(z, t) a_2(z, t) a_3(z, t) \prod_{i=1}^n z_i^{v_i},$$

where

$$\begin{aligned} a_1(z, t) = \exp \left\{ \sum_{i=1}^n p_{vi} z_i \sum_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right\} = \prod_{i=1}^n \prod_{j=1}^n \exp \{ p_{vi} z_i (Y_{ij}(t) - Y_{ij}(0)) \} = \\ = \prod_{i=1}^n \prod_{j=1}^n \sum_{l_i=0}^{\infty} \frac{[p_{vi} z_i (Y_{ij}(t) - Y_{ij}(0))]^{l_i}}{l_i!} = \sum_{l_1=0}^{\infty} \dots \sum_{l_n=0}^{\infty} \frac{p_{v1}^{l_1} \dots p_{vn}^{l_n} z_1^{l_1} \dots z_n^{l_n}}{l_1! l_2! \dots l_n!} \prod_{i=1}^n \prod_{j=1}^n [(Y_{ij}(t) - Y_{ij}(0))]^{l_i}. \\ a_2(z, t) = \exp \left\{ \sum_{i=1}^n \frac{1}{z_i} (A_i(t) - A_i(0)) \right\} = \prod_{i=1}^n \exp \left\{ \frac{1}{z_i} (A_i(t) - A_i(0)) \right\} = \\ = \prod_{i=1}^n \sum_{q_i=0}^{\infty} \frac{[(A_i(t) - A_i(0)) z_i^{-1}]^{q_i}}{q_i!} = \sum_{q_1=0}^{\infty} \dots \sum_{q_n=0}^{\infty} \prod_{i=1}^n \frac{[A_i(t) - A_i(0)]^{q_i}}{q_i!} z_1^{-q_1} \dots z_n^{-q_n}. \end{aligned}$$

$$\begin{aligned}
\alpha_i(z, t) &= \exp \left\{ \sum_{i=1}^n \sum_{j=1}^n \frac{z_i}{z_j} (B_{ij}(t) - B_{ij}(0)) \right\} = \\
&= \prod_{i=1}^n \prod_{j=1}^n \exp \left\{ \frac{z_i}{z_j} (B_{ij}(t) - B_{ij}(0)) \right\} = \prod_{i=1}^n \prod_{j=1}^n \sum_{r_i=0}^{\infty} \frac{[(B_{ij}(t) - B_{ij}(0)) z_i z_j^{-1}]^{r_i}}{r_i!} = \\
&= \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \prod_{i=1}^n \prod_{j=1}^n \frac{[(B_{ij}(t) - B_{ij}(0)) z_i z_j^{-1}]^{r_i}}{r_i!} = \prod_{i=1}^n \prod_{j=1}^n \exp \left\{ \frac{z_i}{z_j} (B_{ij}(t) - B_{ij}(0)) \right\} = \\
&= \prod_{i=1}^n \prod_{j=1}^n \sum_{r_i=0}^{\infty} \frac{[(B_{ij}(t) - B_{ij}(0)) z_i z_j^{-1}]^{r_i}}{r_i!} = \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \prod_{i=1}^n \prod_{j=1}^n \frac{[(B_{ij}(t) - B_{ij}(0)) z_i z_j^{-1}]^{r_i}}{r_i!} = \\
&= \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \frac{\prod_{i=1}^n \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i}}{r_1! \dots r_n!} z_1^{r_1 + r_2 + \dots + r_n} z_2^{r_1 + r_2 + \dots + r_n} \dots z_n^{r_1 + r_2 + \dots + r_n} z_1^{-r_1} z_2^{-r_2} \dots z_n^{-r_n} = \\
&= \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \frac{\prod_{i=1}^n \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i}}{r_1! \dots r_n!} z_1^{R-r_1} \dots z_n^{R-r_n}
\end{aligned}$$

Multiplying $\alpha_0(t)$, $\alpha_1(z, t)$, $\alpha_2(z, t)$, $\alpha_3(z, t)$ and $\prod_{l=1}^n z_l^{y_l}$ obtain the expression for the generating function (7) has the form:

$$\begin{aligned}
\Psi_n(z, t) &= \alpha_0(t) \sum_{l_1=0}^{\infty} \dots \sum_{l_n=0}^{\infty} \sum_{q_1=0}^{\infty} \dots \sum_{q_n=0}^{\infty} \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{l_i! q_i! r_i!} \times \right. \\
&\times \left. \left(\prod_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right)^{l_i} (A_i(t) - A_i(0))^{q_i} \times \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i} z_i^{x_i + l_i - q_i - r_i + R} \right], \quad (13)
\end{aligned}$$

where $R = \sum_{i=1}^n r_i$.

4. Finding the average number of messages in the network systems

It is known that the expectation of the i -th component of a multidimensional random variable can be found by differentiating the generating function (13) and putting on $z_i = 1$, $i = 1, n$. Therefore, the average number of messages in system S_c will use the relation:

$$\begin{aligned}
\frac{\partial \Psi_n(z, t)}{\partial z_c} &= a_0(t) \sum_{l_1=0}^c \dots \sum_{l_n=0}^c \sum_{q_1=0}^c \dots \sum_{q_n=0}^c \sum_{r_1=0}^c \dots \sum_{r_n=0}^c (x_c + l_c - q_c - r_c + R) \times \\
&\times \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{l_i! q_i! r_i!} \left(\prod_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right)^{l_i} (A_i(t) - A_i(0))^{q_i} \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i} \right] \times \\
&\times Z_c^{x_c - l_c - q_c - r_c - R - 1} \prod_{\substack{i=1, \\ i \neq c}}^n Z_i^{x_i + l_i - q_i - r_i + R}.
\end{aligned}$$

It follows that the average number of applications in the system S_c is determined by the formula

$$\begin{aligned}
N_c(t) &= \left. \frac{\partial \Psi_n(z, t)}{\partial z_c} \right|_{z=(1, \dots, 1)} = a_0(t) \sum_{l_1=0}^c \dots \sum_{l_n=0}^c \sum_{q_1=0}^c \dots \sum_{q_n=0}^c \sum_{r_1=0}^c \dots \sum_{r_n=0}^c (x_c + l_c - q_c - r_c + R) \times \\
&\times \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{l_i! q_i! r_i!} \left(\prod_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right)^{l_i} (A_i(t) - A_i(0))^{q_i} \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i} \right]. \quad (14)
\end{aligned}$$

We make in (14) change of variables $k_c = x_c + l_c - q_c - r_c - R$, then $l_c = k_c - x_c + q_c + r_c - R$ and

$$\begin{aligned}
N_c(t) &= a_0(t) \sum_{q_1=0}^c \dots \sum_{q_n=0}^c \sum_{r_1=0}^c \dots \sum_{r_n=0}^c \sum_{k_1=x_1-q_1-r_1-R}^c \dots \sum_{k_n=x_n-q_n-r_n-R}^c k_c \times \\
&\times \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{(k_i - x_i + q_i + r_i - R)! q_i! r_i!} \left(\prod_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right)^{k_i - x_i + q_i + r_i - R} \times \right. \\
&\quad \left. \times (A_i(t) - A_i(0))^{q_i} \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i} \right].
\end{aligned}$$

Because the network operating system under high load mode $k_i = x_i - q_i - r_i + R \geq 1$ and, therefore, $q_i \leq x_i - r_i + R - 1$, thus

$$\begin{aligned}
N_c(t) &= a_0(t) \sum_{k_1=1}^c \dots \sum_{k_n=1}^c \sum_{q_1=0}^c \dots \sum_{q_n=0}^c \sum_{r_1=0}^{x_1-q_1-R-1} \dots \sum_{r_n=0}^{x_n-q_n-R-1} \times \\
&\times \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{(k_i - x_i + q_i + r_i - R)! q_i! r_i!} \left(\prod_{j=1}^n (Y_{ij}(t) - Y_{ij}(0)) \right)^{k_i - x_i + q_i + r_i - R} \times \right. \\
&\quad \left. \times (A_i(t) - A_i(0))^{q_i} \left(\prod_{j=1}^n (B_{ij}(t) - B_{ij}(0)) \right)^{r_i} \right]. \quad (15)
\end{aligned}$$

5. Example 1

Let the intensity $\lambda(t) = \lambda t$, $\mu_i(t) = \mu_i [\cos(\omega_i t) + 1]$, $i = \overline{1, n}$. In this case $\Lambda(t) = \frac{\lambda}{2} t^2$, $\Lambda(0) = 0$, $M_i(t) = \mu_i \left[\frac{\sin(\omega_i t)}{\omega_i} + t \right]$, $M_i(0) = 0$, $i = \overline{1, n}$. Suppose, that the probability of adherence message to the queue at time t is given by $f^{(n)}(t) = 1 - e^{-t}$. From (10) follows that

$$\begin{aligned}\Phi_i(t) &= \int \lambda(t) \varphi_i(t) dt = \int \varphi_i(t) d\Lambda(t) = \varphi_i(t) \Lambda(t) - \int \varphi_i'(t) \Lambda(t) dt, \\ Y_{ij}(t) &= \int \lambda(t) \psi_{ij}(t) dt = \int \psi_{ij}(t) d\Lambda(t) = \psi_{ij}(t) \Lambda(t) - \int \psi_{ij}'(t) \Lambda(t) dt, \\ A_i(t) &= \int \mu_i(t) \alpha_i(t) dt = \int \alpha_i(t) dM_i(t) = \alpha_i(t) M_i(t) - \int \alpha_i'(t) M_i(t) dt, \\ B_{ij}(t) &= \int \mu_i(t) \beta_{ij}(t) dt = \int \beta_{ij}(t) dM_i(t) = \beta_{ij}(t) M_i(t) - \int \beta_{ij}'(t) M_i(t) dt,\end{aligned}$$

and we get that

$$\begin{aligned}a_{ij}(t) &= \exp \left\{ - \sum_{i=1}^n \left[\frac{\lambda}{2} (t^2 - \varphi_i(t) t^2 + \int \varphi_i'(t) t^2 dt) p_{\varphi_i} + \right. \right. \\ &\quad \left. \left. + \mu_i \left(\frac{\sin(\omega_i t)}{\omega_i} + t - \beta_{ij}(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) + \int \beta_{ij}'(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) dt \right) \right] \right\} - \\ &= \exp \left\{ - \sum_{i=1}^n \left[\frac{\lambda}{2} (t^2 - \int \varphi_i(t) t dt) p_{\varphi_i} + \mu_i \left(\frac{\sin(\omega_i t)}{\omega_i} + t - \int \beta_{ij}(t) (\cos(\omega_i t) + 1) dt \right) \right] \right\}.\end{aligned}$$

$\Phi_i(0) = 0$, $Y_{ij}(0) = 0$, $\Lambda(0) = 0$, $B_{ij}(0) = B_{ij}(0) = 0$, $B_{ij}(0) = 0$, $i, j = \overline{1, n}$. Conditional probabilities $\varphi_i(t)$, $\psi_{ij}(t)$, $\alpha_i(t)$ and $\beta_{ij}(t)$, according to (1)-(3), found from the relations

$$\begin{aligned}\varphi_i(t) &= (1 - f^{(n)}(t)) \left(p_{\varphi_i} + \sum_{j=1}^n p_{\psi_j} \varphi_j(t) \right), \quad \psi_{ij}(t) = f^{(n)}(t) \delta_{ij} + (1 - f^{(n)}(t)) \sum_{j=1}^n p_{\psi_j} \psi_{ij}(t), \\ \alpha_i(t) &= p_{\alpha_i} + \sum_{j=1}^n p_{\alpha_j} \alpha_j(t), \quad \beta_{ij}(t) = \sum_{j=1}^n p_{\beta_j} \beta_{ij}(t), \quad i, j = \overline{1, n}.\end{aligned}\quad (16)$$

Solving the systems of linear equations (16) in the package *Mathematica*, analytical solutions can be obtained, but they are cumbersome already at $n = 3$. For example, expression for the conditional probability of a time-dependent $\varphi_i(t)$, at $n = 3$ has the form:

$$\begin{aligned} \varphi_{13}(t) = & \left((e^{-t} p_{13} (e^{-2t} p_{22} - 1) - e^{-3t} p_{12} p_{23}) (e^{-4t} p_{13} p_{30} - e^{-t} p_{10} (e^{-3t} p_{33} - 1)) - \right. \\ & \left. - (e^{-3t} p_{13} p_{20} - e^{-3t} p_{10} p_{23}) (e^{-4t} p_{13} p_{32} - e^{-t} p_{12} (e^{-3t} p_{33} - 1)) \right) / \\ & / \left((e^{-t} p_{13} (e^{-3t} p_{22} - 1) - e^{-3t} p_{12} p_{23}) (e^{-4t} p_{13} p_{31} - (e^{-t} p_{11} - 1) e^{-t} (e^{-3t} p_{33} - 1)) - \right. \\ & \left. - (e^{-3t} p_{13} p_{21} - e^{-2t} (e^{-t} p_{11} - 1) p_{23}) (e^{-4t} p_{13} p_{32} - e^{-t} p_{12} (e^{-3t} p_{33} - 1)) \right), \end{aligned}$$

and for the probability $\psi_{21}(t)$:

$$\begin{aligned} \psi_{21}(t) = & (e^{4t} p_{20} - e^{3t} p_{11} p_{20} - e^{3t} p_{21} + e^{2t} p_{21} + e^{3t} p_{10} p_{21} + p_{13} p_{21} p_{30} + e^t p_{23} p_{30} - \\ & - p_{11} p_{23} p_{30} - p_{13} p_{20} p_{31} - p_{23} p_{31} + e^t p_{23} p_{31} + p_{10} p_{23} p_{31} - e^t p_{20} p_{33} + p_{21} p_{20} p_{33} + \\ & + p_{21} p_{33} - e^t p_{21} p_{33} - p_{10} p_{21} p_{33}) / (e^{6t} + e^{5t} p_{11} - \\ & + e^{3t} p_{12} p_{21} + e^{3t} p_{22} - e^{3t} p_{11} p_{22} + e^{2t} p_{13} p_{31} - p_{13} p_{22} p_{31} + p_{12} p_{23} p_{31} - p_{13} p_{21} p_{32} + \\ & + e^t p_{23} p_{32} - p_{11} p_{23} p_{32} + e^{3t} p_{33} - e^{2t} p_{11} p_{33} - p_{12} p_{21} p_{33} - e^t p_{22} p_{33} - p_{11} p_{22} p_{33}). \end{aligned}$$

Expression (13) takes the form

$$\begin{aligned} \Psi_n(z, t) = & a_0(t) \sum_{l_1=0}^{\infty} \dots \sum_{l_n=0}^{\infty} \sum_{q_1=0}^{\infty} \dots \sum_{q_n=0}^{\infty} \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \prod_{i=1}^n \left[\frac{p_{0i}^{l_i}}{l_i! q_i! r_i!} \times \right. \\ & \times \left(\prod_{j=1}^n \left(\frac{\lambda}{2} (\psi_{ij}(t) t^2 - \int \psi'_{ij}(t) t^2 dt) \right) \right)^{l_i} \times \\ & \times \left(\mu_i \left(\alpha_i(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) - \int \alpha'_i(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) dt \right) \right)^{q_i} \times \\ & \times \left(\mu_i \prod_{j=1}^n \left(\beta_{ij}(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) - \int \beta'_{ij}(t) \left(\frac{\sin(\omega_i t)}{\omega_i} + t \right) dt \right) \right)^{r_i} z_i^{x_i + l_i - q_i - r_i + R} \Big], \end{aligned}$$

after simplification we obtain:

$$\begin{aligned} \Psi_n(z, t) = & a_n(t) \sum_{l_1=0}^{\infty} \dots \sum_{l_n=0}^{\infty} \sum_{q_1=0}^{\infty} \dots \sum_{q_n=0}^{\infty} \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \prod_{i=1}^n \left[\left(\frac{\lambda}{2} t^2 \right)^{l_i} \frac{\mu_i^{q_i + r_i} p_{0i}^{l_i}}{l_i! q_i! r_i!} \times \right. \\ & \times \left(\prod_{j=1}^n \int \psi_{ij}(t) t dt \right)^{l_i} \left(\int \alpha_i(t) (\cos(\omega_i t) + 1) dt \right)^{q_i} \times \\ & \left. \left(\prod_{j=1}^n \int \beta_{ij}(t) (\cos(\omega_i t) + 1) dt \right)^{r_i} z_i^{x_i + l_i - q_i - r_i + R} \right] \quad (17) \end{aligned}$$

Suppose we have found, for example, the probability of state It is the coefficient of $z_1 z_2 \dots z_n$ in the expansion of $\Psi_n(z, t)$ in multiple series (17), so that when the degree of z_i must satisfy the relation $x_i + l_i - q_i - r_i + R = 1$, $i = \overline{1, n}$, it follows that

$$q_i = x_i + l_i + \sum_{j=i}^{i-1} r_j - 1, \quad i = \overline{1, n}, \quad q_i + r_i = x_i + l_i + \sum_{j=i}^{i-1} r_j - 1, \quad i = \overline{1, n},$$

$$l_i + q_i + r_i = x_i + 2l_i + \sum_{j=i}^{i-1} r_j - 1, \quad i = \overline{1, n}, \quad \sum_{i=1}^n (l_i + q_i + r_i) = \sum_{i=1}^n (x_i + 2l_i) + n(R-1).$$

Therefore, from (17) follows that

$$P(1,1,\dots,1,t) = a_0(t) \sum_{l_1=0}^{\infty} \dots \sum_{l_n=0}^{\infty} \sum_{r_1=0}^{\infty} \dots \sum_{r_n=0}^{\infty} \prod_{i=1}^n \left[\frac{P_{0i}^{l_i}}{l_i! \left(x_i + l_i + \sum_{j=i}^{i-1} r_j - 1 \right)! r_i!} \left(\frac{\lambda}{2} t^2 \right)^{\sum_{i=1}^n l_i} \mu_i^{x_i + l_i + R-1} \times \right. \\ \left. \times \left(\prod_{i=1}^n \int \psi_{ii}(t) dt \right)^{l_i} \left(\int \alpha_i(t) (\cos(\omega_i t) + 1) dt \right)^{x_i + l_i} \prod_{i=1}^n \left(\int \beta_{ii}(t) (\cos(\omega_i t) + 1) dt \right)^{r_i} \right].$$

We assume that $n=10$ and the initial time the network is in a state $(2,2,2,2,2,2,2,2,2,0)$. We find the probability of the state $P(1,1,1,1,1,1,1,1,1,1,t)$ using the formula (17). Let the intensities $\lambda(t) = \lambda t$, $\mu_i(t) = \mu_i [\cos(\omega_i t) + 1]$, transition probabilities of messages are equal: $p_{iv} = 0.1, i = \overline{1,10}$, $p_{vi} = 0.1, i = \overline{0,9}$, $p_{i0} = 0.5, i = \overline{1,9}$, $p_{i10} = 0.5, i = \overline{1,9}$, $p_{00} = 0, i = \overline{0,10}$. In addition, the intensity of the service messages in the systems are: $\mu_i = 20.8, i = \overline{1,4}$, $\mu_i = 10.22, i = \overline{4,9}$, $\mu_{10} = 20.5$. The expression for the time-dependent probability of the state in the systems of the network obtained by on a computer using a mathematical calculation package *Mathematica*. Figure 1 shows a chart of this probability depending on the time t .

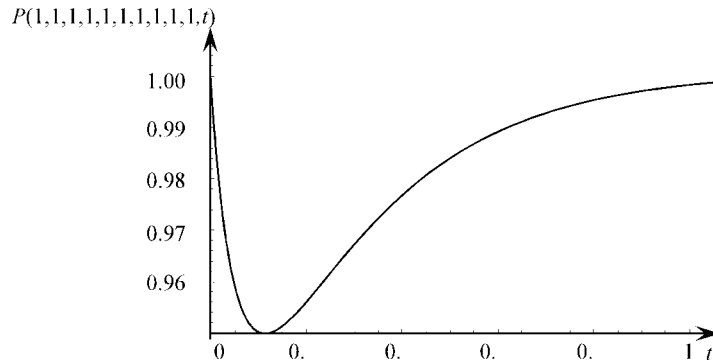


Fig. 1. The chart of probability of the state $P(1,1,1,1,1,1,1,1,1,1,t)$

6. Example 2

We will consider the network described in Example 1. Let $\lambda(t) = \lambda t$, $\mu_i(t) = \mu_i [\cos(\omega_i t) + 1]$, transition probabilities of messages are equal: $p_{0i} = 0.1, i = 1, 10$, $p_{10i} = 0.1, i = 0, 9$, $p_{ri} = 0.5, i = 1, 9$, $p_{r10} = 0.5, i = 1, 9$, $p_{ii} = 0, i = 0, 10$, intensity of service messages in the systems are: $\mu_i = 20.8, i = 1, 4$, $\mu_i = 10.22, i = 4, 9$, $\mu_{10} = 20.5$. The average number of messages in the systems network (in the queue and service) at the initial time $t = 0$ equally $N_i(0) = 0, i = 1, 10$. Equation (15) to find the average number of messages in the systems network when the network parameters and the conditional probabilities depend on the time takes the form:

$$N_i(t) = \exp \left\{ - \sum_{j=1}^n \left[\frac{\lambda}{2} (t^2 - \int \varphi_j(t) t dt) p_{0i} + \right. \right. \\ \left. \left. + \mu_i \left(\frac{\sin(\omega_j t)}{\omega_j} + t - \int \beta_{ij}(t) (\cos(\omega_j t) + 1) dt \right) \right] \right\} \times \sum_{k_1=1}^{\epsilon} \dots \sum_{k_n=1}^{\epsilon} k_i \sum_{r_1=0}^{\epsilon} \dots \sum_{r_n=0}^{\epsilon} \sum_{q_1=0}^{x_1 - r_1 - R - 1} \dots \sum_{q_n=0}^{x_n - r_n - R - 1} \times \\ \times \prod_{j=1}^n \left[\frac{\lambda^{k_i - v_i - q_i + r_i - R} p_{0i}^{r_i}}{(k_i - x_i + q_i + r_i - R)! q_i! r_i!} \left(\prod_{j=1}^n \int \psi_{ij}(t) t dt \right)^{k_i - v_i - q_i - r_i - R} \times \right. \\ \left. \times \left(\int \alpha_i(t) (\cos(\omega_i t) + 1) dt \right)^{r_i} \left(\prod_{j=1}^n \int \beta_{ij}(t) (\cos(\omega_j t) + 1) dt \right)^{r_i} \right].$$

This example is designed on a computer using a mathematical calculation package *Mathematica*. Here are some of the values of the average number of applications in the systems of the network (in the queue and service) at time $t = 0.2$, found using the program: $N_1(t) = 0.547$, $N_2(t) = 0.323$, $N_3(t) = 0.429$, $N_4(t) = 0.522$, $N_5(t) = 0.742$, $N_{10}(t) = 0.654$. Figure 2 shows a chart of the average number of messages in the QS S_3 depending on time t .

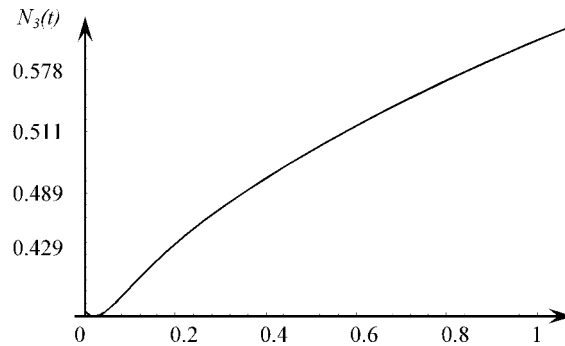


Fig. 2. The chart of changes of the average number of messages $N_i(t)$ in QS S_i

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